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Synthesizing a network of evidence on a seabird bycatch mitigation measure

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5 **Synthesizing a network of evidence on a seabird bycatch mitigation measure**

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20 **ABSTRACT**

21 Robust estimates of the relative efficacies of alternative management interventions are essential
22 for developing evidence-informed fisheries bycatch policy. Bycatch is a major threat to the
23 conservation of albatrosses and other pelagic seabirds. Branchline weighting is one approach
24 prescribed by regional fisheries management organizations and the Agreement on the
25 Conservation of Albatrosses and Petrels to reduce seabird bycatch in pelagic longline fisheries.
26 We used a Bayesian multilevel network meta-regression modelling approach to conduct the first
27 synthesis of available evidence to assess the relative efficacies at mitigating seabird bycatch of
28 alternative pelagic longline weighting designs. Unlike conventional pairwise meta-analysis,
29 network meta-analysis enables the simultaneous comparison of multiple interventions within a
30 coherent modelling framework. There was a >97% probability that all weighting designs
31 significantly reduced seabird bycatch compared to a reference design with no weight within 5m
32 of the hook. Nonetheless, some weighting designs were significantly more effective at reducing
33 seabird bycatch than others — for instance, the 2 designs with weights >60g and >1m from the
34 hook performed the best with >93% probability that those 2 designs performed significantly
35 better than 2 more commonly used designs with less weight but attached closer to the hook.
36 These two best performing designs reduced seabird bycatch by *ca.* 89% relative to the
37 reference design. These relative efficacies and rankings, when combined with other
38 performance criteria such as costs to commercial viability and crew safety, support robust
39 evaluations of alternative bycatch management strategies.

40
41 **Keywords:** Bayesian network meta-analysis; Bycatch mitigation; Evidence synthesis; Fishing
42 gear weights; RFMO; Seabirds

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45 **1. INTRODUCTION**

46 Incidental capture in fisheries is a major threat to the conservation of pelagic seabirds and
47 reduces fishing efficiency (Brothers et al., 1999; Anderson et al. 2011; Phillips et al., 2016). This
48 mortality source is particularly problematic for albatrosses and petrels, which are two of the
49 three most threatened groups of seabirds (Phillips et al. 2016; Dias et al. 2019). Of the suite of
50 anthropogenic threats that albatrosses and petrels are exposed to, fishing-associated mortality
51 causes the largest population impacts (Heppell et al., 2000; Bakker et al., 2018). For depleted

52 populations, even low fishing-associated mortality levels can have profound effects and risk
53 extirpation (Heppell et al. 2000).

54 Pelagic longlining is a main fishing gear type supplying canned, frozen and fresh
55 products from tuna and tuna like species (Scombroidei), billfishes (Xiphoidei) and other large
56 pelagic and neritic apex- and meso-predators (Miyake et al., 2010; Williams and Ruaia, 2022).
57 Primarily at higher latitudes, pelagic longline fisheries are a major hazard to exposed pelagic
58 seabirds (Anderson et al. 2011; Dias et al. 2019). Several methods can substantially reduce
59 seabird bycatch rates in pelagic longline fisheries that cause minimal or no costs to economic
60 viability, practicality and crew safety (Hall et al., 2017; ACAP 2023). However, there has been
61 mixed progress in the adoption of seabird bycatch mitigation measures by fisheries
62 management authorities and compliance is uncertain due to deficits in fisheries monitoring,
63 surveillance and enforcement (Juan-Jorda et al., 2018).

64 Branchline weighting is one option for seabird bycatch mitigation in pelagic longline
65 fisheries, which is typically prescribed to be used in combination with other seabird bycatch
66 mitigation methods. Some pelagic longline vessels conventionally incorporate lead-centered
67 weighted swivels into branchlines of various mass and locations to: increase the gear sink rate;
68 maintain the gear at a desired fishing depth; affect the action (movement) of the terminal baited
69 hooks; and reduce line tangles (Brothers et al., 1999; Beverly et al., 2003). The mass of pelagic
70 longline branchline weights and distance between the weight and hook significantly affect
71 seabird catch risk during setting (Robertson et al., 2013; Melvin et al., 2013, 2014; Santos et al.,
72 2019), as well as during the gear soak and haul (Gilman et al., 2014b; Phillips and Wood, 2020).
73 During setting, these two variables affect the sink rate of baited hooks and their availability to
74 seabirds. The initial baited hook sink rate affects seabird catch risk near the surface. The sink
75 rate of a baited hook begins to be affected by a branchline weight once the weight sinks to a
76 depth when the leader (the section of the branchline between the hook and weight) becomes
77 taut (Robertson et al., 2010). The closer a weight is to the hook and the larger the mass, the
78 less time it takes for the weight to affect the hook's sink rate near the surface and the less time
79 baited hooks are available to seabirds to the extent of their diving depth (Robertson et al., 2013;
80 Jimenez et al., 2019). Similarly, during the gear soak, in fisheries where hooks fish near the
81 surface, in deeper-set fisheries when catch bring the gear to the surface, as well as when the
82 gear shoals from currents and wind, and during the haul when hooks are brought back to the
83 surface, the closer a weight is to the hook, the more likely the weight will reduce the availability
84 of baited hooks to surface-foraging seabirds (Gilman et al., 2014b).

85 In regions where secondary seabird interactions occur or where primary interactions
86 occasionally occur at depth, the hook sink rate to depths attainable by these deep-diving
87 species is an additional important predictor of seabird catch risk (Jimenez et al., 2012; Melvin et
88 al., 2014; Rollinson et al., 2016; Guilford et al., 2022). Secondary interactions occur when deep-
89 diving seabirds access baited hooks at depth and return them to the sea surface, providing
90 larger primarily surface-foraging seabird species with a second opportunity to access baited
91 hooks (Jimenez et al., 2012; Melvin et al., 2014). For fisheries in these regions, the effect of the
92 branchline weighting design on the baited hook sink rate both near the surface where most or all
93 primary interactions occur and to as deep as *ca.* 25m for secondary interactions by proficient
94 divers such as species of shearwaters and *Procellaria* petrels (Bentley et al., 2021; Rollinson et
95 al., 2016; Frankish et al., 2021) will affect seabird catch risk. In some regions, such as the
96 central north Pacific Ocean, secondary interactions are rare events and seabirds susceptible to
97 capture typically only make body thrusts to reach prey within *ca.* 1 m of the surface (Prince et
98 al., 1994; Kazama et al., 2019; Bentley et al., 2021, Guilford et al., 2022). For fisheries in these
99 regions, the effect of the branchline weighting design on the sink rate very near the surface has
100 a greater effect on seabird bycatch risk than the sink rate to depth.

101 Unlike many other seabird bycatch mitigation methods, robust compliance monitoring of
102 a required branchline weighting design when conventional lead-centered swivels are used,

103 which are crimped in place, is feasible by dockside inspections for vessels making relatively
104 short trips, where it would be impractical for crew to move weights during a fishing trip (ACAP,
105 2023). This is an important consideration given the extremely limited onboard monitoring that
106 occurs in global pelagic longline fisheries (Gilman et al., 2014a). Surveillance of line weighting
107 designs of vessels making long trips would, however, require onboard observers (and possibly
108 by electronic monitoring systems, Gilman et al., 2019). The location of attachment of sliding
109 weights and weighted hook shielding devices can be easily changed by crew during both short
110 and long trips, making dockside inspections an ineffective compliance monitoring approach.
111 This could potentially be addressed by requiring unweighted swivels to be attached at the
112 maximum permitted distance that weights can be attached from hooks.

113 Five regional fisheries management organizations (RFMOs) have adopted binding
114 measures that include branchline weighting as one option to meet seabird bycatch mitigation
115 requirements by pelagic longline fisheries in designated areas, where each RFMO defines three
116 or four options for weighting designs (IATTC, 2011; ICCAT, 2011; WCPFC, 2018; SIOFA, 2022;
117 IOTC, 2023). The Agreement on the Conservation of Albatrosses and Petrels (ACAP)
118 recommends three weighting design options (ACAP, 2023). ACAP's (2023) recommendations
119 are based on mechanistic studies that measured sink rates of alternative branchline weighting
120 designs when considering whether the estimated sink rates are predicted to afford adequate
121 protection when a vessel is using a bird-scaring *tori* line and considering seabird diving depths.
122 The ACAP (2023) recommendations are also based on findings from three selected studies on
123 the effects of alternative weighting designs on seabird interaction rates (Gianuca et al. 2011;
124 Barrington et al., 2016; Jiménez et al. 2019a; Santos et al. 2019).

125 The controlled experiments have a higher strength of evidence compared to sink rate
126 mechanistic studies, while findings from meta-analytic synthesis studies have the highest
127 strength of evidence, producing the most robust and generalizable findings that are optimal for
128 informing global and regional seabird bycatch policy (Gilman and Chaloupka, 2023). A synthesis
129 of accumulated evidence of seabird bycatch rate responses to alternative branchline weighting
130 designs, including those prescribed by RFMOs and ACAP, has not previously been conducted.
131 Meta-analytic syntheses, which synthesize estimates sourced from a mixture of independent
132 and context-specific studies, have the lowest risk of error and bias and the highest strength of
133 evidence, producing the most robust and generalizable findings that are optimal for evaluating
134 alternative regional bycatch management strategies (Dicks et al. 2014; Nichols et al. 2019).

135 This study provides the strongest evidence to guide regional policies on branchline
136 weighting to mitigate seabird bycatch. The study objective was to determine whether certain
137 pelagic longline branchline weighting designs, including those required by RFMOs and
138 recommended by ACAP, have similar or otherwise significantly different seabird catch rates. We
139 also explored informative predictors of seabird bycatch risk other than branchline weighting. We
140 conducted a systematic literature review to compile relevant publications and assembled a
141 dataset suitable for meta-analytic evaluation. We used a Bayesian multilevel network meta-
142 regression modelling approach to synthesize available evidence to assess the relative efficacies
143 of branchline weighting designs that are used in pelagic longline fisheries to mitigate seabird
144 bycatch. A key outcome of the network meta-synthesis is a league table or rankings of those
145 relative efficacies that enable identifying the least to most efficacious designs. Findings can be
146 used by fisheries management authorities to define evidence-informed weighting designs that
147 meet objectives for managing seabird bycatch in pelagic longline fisheries.

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150 2. METHODS

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2.1. Data Compilation

153 A two-tiered systematic and unstructured literature search was employed to compile relevant
154 peer-reviewed published and grey literature, including government observer program data, and
155 to assemble a dataset suitable for meta-analytic evaluation. The methods for the systematic
156 literature review were adapted from the Reporting Standards for Systematic Evidence
157 Syntheses (Haddaway et al., 2018), Collaboration for Environmental Evidence (Pullin et al.,
158 2020), and extended PRISMA framework applicable for network meta-analyses (NMA) (Hutton
159 et al 2015). Details on the literature review methods, including searched databases, search
160 strings, study eligibility criteria and screening process for retention, and extracted data fields are
161 in Supplemental Material Section S1. There were 12 retained publications obtained from the
162 systematic review (Supplemental Material Fig. S1, Table S2). One additional retained
163 publication (Table S2) and five retained observer program datasets from New Zealand, Taiwan,
164 the U.S. Atlantic Ocean, the Hawaii-based U.S. Pacific Ocean shallow-set swordfish fishery,
165 and the Hawaii-based U.S. Pacific Ocean deep-set tuna fishery (Supplemental Material Section
166 S4) were obtained from the unstructured literature review.

167 168 **2.2. Branchline Weighting Design Treatments**

169 Supplemental Material Table S1a summarizes the branchline weighting designs prescribed by
170 five RFMOs with binding seabird measures and recommended by ACAP. To produce a
171 manageable number of weighting design treatments and with adequate sample sizes we used
172 only the \geq designs (e.g., the prescribed design $\geq 45\text{g}$ weights within 1m of the hook was retained
173 and eliminated $>45\text{g}$ within 1m). As none of the compiled records had >1 weight within half a
174 meter of the hook, we combined the design prescribed by one of the RFMOs that specifies this
175 limit with an ACAP recommended design of $\geq 40\text{g}$ within 0.5m. We adapted the RFMO/ACAP
176 designs to provide more detailed definitions so that a single branchline weighting design can fall
177 under only one unique category. An RFMO-prescribed designs of $\geq 98\text{g}$ attached between
178 $>3.5\text{m}$ and $\leq 4\text{m}$ of the hook was excluded because there was only a single record in this
179 category. We included two categories not prescribed by RFMOs or ACAP. This resulted in the
180 following 7-category treatment variable for branchline weighting design for the modeling, where
181 A is referred to as the reference design and is not prescribed by RFMOs or ACAP, designs B
182 through F are prescribed RFMO or ACAP options (sample sizes are number of records), and
183 design G is not prescribed by RFMOs or ACAP:

- 184
- 185 A. 0g attached $\leq 5\text{m}$ from the hook (N=86)
- 186 B. $\geq 40\text{g}$ attached $\leq 0.5\text{m}$ from the hook (N=42)
- 187 C. $\geq 45\text{g}$ to $<60\text{g}$ between $>0.5\text{m}$ and $\leq 1\text{m}$ from the hook (N=37)
- 188 D. $\geq 60\text{g}$ between $>0.5\text{m}$ and $\leq 1\text{m}$ from the hook (N=27)
- 189 E. $\geq 60\text{g}$ to $<80\text{g}$ between $>1\text{m}$ and $\leq 3.5\text{m}$ or $\geq 80\text{g}$ between $>2\text{m}$ and $\leq 3.5\text{m}$ from the hook
190 (N=18)
- 191 F. $\geq 80\text{g}$ between $>1\text{m}$ and $\leq 2\text{m}$ from the hook (N=9)
- 192 G. $\geq 8\text{g}$ to $\leq 39\text{g}$ between $\leq 1\text{m}$ or $\geq 45\text{g}$ to $\leq 80\text{g}$ between $\geq 1.3\text{m}$ and $\leq 4\text{m}$ from the hook
193 (N=52)
- 194

195 Table S1b identifies which of the combined RFMO and ACAP design options are met by each of
196 the above 7 branchline weighting design treatments included in the NMA. Table S3 summarizes
197 the records covered by each of the 7 treatments, including the number of fisheries and regions
198 from which they were derived, and the range and median weight mass and leader lengths of the
199 records within each treatment.

200 201 **2.3. Statistical Modelling Approach**

202 203 **2.3.1. Visualising the network geometry**

204 We summarise the underlying NMA structure for our study in Figure 1, which comprises the 7
205 pelagic longline branchline weighting designs (A-G) with 21 weighting design combinations or
206 interactions — this plot is a mandatory component of any network modelling workflow (Liu et al
207 2023). Recently, Freeman et al (2023) also proposed the use of upset plots (Conway et al 2017)
208 for further exploring and visualising the NMA dataset characteristics, which we implemented
209 using the `ComplexUpset` R package (Krassowski 2020) for the weighting design combinations
210 shown in the network geometry plot (Figure 1). The upset plot display of the network structure is
211 shown in Figure S2.

213 **2.3.2. Statistical modelling framework**

214 NMA is a well-established statistical modelling approach used in the healthcare evaluation,
215 pharmacometrics, psychology and clinical epidemiology domains for comparing the relative
216 efficacies of three or more interventions or treatments by synthesising aggregate and/or individual
217 participant evidence sourced from multiple trials (Ades et al 2024). An accessible introduction to
218 Bayesian NMA can be found in Liu et al (2023), the comparative benefits of Bayesian inference-
219 based meta-analysis can be found in (Pappalardo et al 2020), while an authoritative primer on
220 Bayesian statistical modelling in general can be found in van de Schoot et al (2021).

221 Conventional pairwise meta-analysis is based on synthesising direct evidence
222 comparing only two treatments from multiple trials that can only generate so-called unrelated
223 treatment effects. For instance, say there are three possible treatments A-C but only AB and AC
224 comparisons are known from direct head-to-head clinical trials — no BC trials have ever
225 occurred and so any inference about such an effect remains unknown in conventional pairwise
226 meta-analysis. NMA leverages both direct and indirect evidence for three or more treatment
227 comparisons, so the BC effect can be inferred indirectly as part of a connected network of trials
228 based on direct evidence from the AB and AC trials — a property known as *coherence* (Ades et
229 al 2024). The relative rankings of those NMA-based treatment effects can then be used to
230 identify the “best” treatment, given that specific set of treatments, and to support evidence-
231 informed cost-effectiveness treatment recommendations given those rankings.

232 Specifically, we used a Bayesian multilevel (or random-effects) network meta-regression
233 modelling approach (Phillippo et al 2020) to synthesis all available studies comparing longline
234 gear weighting designs on seabird bycatch rates (Figure 1). Importantly, not all seven weighting
235 designs need to be in each study so long as each study comprises a subset of at least two or
236 more of the designs. We could only source the study-specific summarised or aggregate data
237 and not the pelagic longline set-specific data for each study, although a mix of both data
238 sources, if available, can be combined in a single integrated Bayesian multilevel network meta-
239 regression model (Phillippo et al 2020, Singh et al 2022).

240 We also sourced a range of study-specific marginal covariate summaries that we could
241 then use as potentially informative seabird bycatch predictors in our network meta-regression
242 modelling framework. See Supplemental Material Sections S1 and S2 for a description of
243 potentially informative predictors that were extracted from retained publications and explored for
244 potential inclusion in the models. The explored predictors included latitude, bait type, hooks
245 between floats, oceanic region and five binary predictors for the following additional seabird
246 bycatch mitigation measures: tori line, night setting, side setting, hook shielding device, and
247 blue-dyed bait (Supplemental Material Table S2).

248 We fitted our multilevel network meta-regression models using the Stan computation
249 back-end (Carpenter et al 2017) via the `multinma` interface for R (Phillippo 2024) with Poisson
250 likelihood (log link) appropriate for study-specific aggregate seabird catch data and with the
251 study-specific aggregate number of hooks (fishing effort) as the exposure metric. The random-
252 effects structure (intercepts only) comprised the study-specific identities. These models are
253 therefore Poisson likelihood GLMMs with the response or effect size being the study-specific log
254 rate ratio (Dias et al 2011, Hooper et al 2000) – also known as the log response ratio

255 (Lajeunesse 2011). We also used QR-decomposition and numeric regression term centering in
256 our network meta-regression models, fitted using `multinma` (Phillippo 2024), to improve
257 computational efficiency by guarding against potential multicollinearity among covariates that
258 might affect the model posterior geometry (Betancourt 2017).

259 All models were fit with 4 Markov chains with 5,000 iterations per chain after a warm-up of
260 1000 iterations. Model convergence was assessed using parameter-specific diagnostics such as
261 multiple chain rank plots, bulk and tail effective sample size metrics and a rank-based *Rhat*
262 statistic (Vehtari et al 2021). All inference was based on the 20,000 network meta-regression
263 model posterior draws, which were accessed as multidimensional arrays to support a range of
264 post-processing functions provided by the `bayesplot` (Gabry et al 2019) and `tidybayes` R
265 packages (Kay 2023). For instance, those posterior draws (samples) were used to derive the
266 uncertainty intervals (HDIs or highest posterior density intervals: Kruschke and Liddell 2018; and
267 see also Pappalardo et al 2020). The `ggplot2` (Wickham 2016), `ggdist` (Kay 2024) and
268 `colorspace` (Zeileis et al 2020) packages for R were used for all the summary graphics while
269 the `patchwork` R package (Pedersen 2020) was used for all multi-panel arrangements.

270

271 **2.3.3. Relative ranking of weighting designs**

272 Using a Bayesian inference framework for our network meta-synthesis of the 7 branchline
273 weighting designs (or treatments) enables us to derive the full posterior probability distribution of
274 all the relative weighting design effects and hence robust summaries of uncertainty of those
275 effects. Additionally, this framework supports the robust ranking of those relative treatment
276 effects in the network, which is especially helpful for supporting evidence-informed bycatch
277 mitigation policy. In a Bayesian modelling framework, the ranks are derived from the effect-
278 specific posterior distributions using two approaches based on: (1) the mean or median of the
279 posterior distribution of the ranks with the 95% credible interval (Phillippo et al 2020), or (2) the
280 probability of a specific treatment ranking the best determined as the Surface Under the
281 Cumulative RAnking curve (SUCRA: Veroniki et al 2018, Daly et al 2019). The SUCRA
282 probabilities are appealing in terms of the ease of its interpretation and Liu et al (2023) propose
283 that rank probabilities > 0.90 reflect an efficacious treatment effect. On the other hand, the
284 posterior ranking approach is more robust than SUCRA to underlying model assumptions. We
285 used both approaches here (posterior median ranking, SUCRA).

286

287 **2.3.4. Model selection using cross-validation**

288 Model selection procedures were based on leave-one-out cross-validation (LOOCv) metrics to
289 estimate any comparative difference in expected predictive accuracy between the various
290 models fitted (Vehtari et al 2017) — such as whether the model included covariates or not. The
291 weight of evidence in favour of one model over any other candidate models was also assessed
292 using Bayesian stacking, which is the Bayesian analogue of model averaging (Yao et al 2018).

293

294 **2.3.5. Assessing network model assumptions**

295 According to Seide et al (2020), there are three fundamental assumptions in NMA: (1)
296 transitivity, (2) network consistency and (3) network heterogeneity (see also Ades et al 2024
297 and the concept of exchangeability). Not all our studies comprised all the 7 weighting designs in
298 head-to-head comparisons (see upset plot, Figure S2), so we assume that any missing designs
299 within a study are missing at random — there are no consensus statistical tests of this specific
300 assumption, but it is conceptually related to the following consistency assumption (Seide et al
301 2020).

302 The all-important consistency assumption concerns whether the direct (head-to-head)
303 estimates and the indirect estimates (made by inference) within each contrast in the connected
304 network are sufficiently similar. This assumption can be readily evaluated using either so-called

305 global assessment tests (unrelated mean effects or UME model: Spinelì 2022) and local
306 assessment tests (node-splitting procedure: Veroniki et al 2013) — we used both forms of
307 consistency assessment tests applied to our best-fit Bayesian network meta-regression model
308 used for inference. The global assessment was done by comparing the expected predictive
309 accuracy between the best-fit model and the unrelated mean effects model using leave-one-out
310 cross-validation (see `Model selection` section above). The UME model is based only on direct
311 evidence and estimates each treatment comparison as a separate and unrelated parameter
312 (Spinelì 2022).

313 The local consistency assumption was specifically assessed using visual review of node-
314 splitting plots (Currier et al 2023) derived using the `multinma` package for R (Phillippo 2024).
315 The heterogeneity assumption is especially important for small-sample studies (Ades et al 2024)
316 and can be readily explore using prior sensitivity analysis, which we implement in our modelling
317 workflow by fitting meta-regression models with weakly informative priors (Lemoine et al 2019)
318 and then using more informative priors recommended for various forms of the between-study
319 heterogeneity parameter (Turner et al 2015, Seide et al 2020).

320

321 **2.3.6. Weighting design-specific relative risk reduction in seabird bycatch**

322 We back-transformed the predicted seabird bycatch log relative risk contrasts derived for the 6
323 branchline weighting design treatments (B-G) relative to reference design A (no weight ≤5m
324 from the hook) using the transform-then-summarize posterior samples approach since such
325 summary metrics are not transform invariant (Wang et al 2018). We then used these back-
326 transformed posterior samples to calculate the mean percent relative reduction in seabird
327 bycatch risk (Gilman et al 2020) for the 6 branchline weighting design treatments relative to the
328 reference design. As in Gilman et al (2020), we also used the highest posterior density interval
329 (HDI) as our measure of uncertainty (Kruschke and Liddell 2018). The mean and 95% HDIs
330 were summarized from the posterior samples for each design relative to the reference design
331 using the `tidybayes` package for R (Kay 2023).

332

333

334 **3. RESULTS**

335

336 **3.1. Model Evaluation Diagnostics**

337 Leave-one-out cross-validation (LOOcv) and Bayesian stacking metrics suggested that the best-
338 fit model fitted to the pelagic longline weighting design network geometry (Figure 1) was a meta-
339 regression model that included predictors for oceanic region and the 5 seabird bycatch
340 mitigation measures (looic difference = -105.9, se = 75.9; ca. 98% weight of evidence in favour
341 of model with covariates) — there was little support for a NMA model without the covariates.
342 Convergence diagnostics such as multiple chain rank plots, and the effective posterior sample
343 size (ESS) metrics coupled with the rank-based diagnostic statistic $Rhat < 1.01$ (Vehtari et al
344 2021), all reflected convergence of the Bayesian network meta-regression model with Poisson
345 likelihood that was used here for statistical inference. Nonetheless, residual deviance for all
346 models fitted here was greater than the number of data points, which can suggest that there
347 might be grounds for further improvement in model fit (Phillippo et al 2020) — see `Covariate
348 effects` section below.

349

350 **3.2. Prior Sensitivity**

351 We assessed comparative predictive accuracy of three between-study heterogeneity parameter
352 priors for the meta-regression model: (1) a weakly informative *half-normal*(scale=5²) prior on the
353 heterogeneity parameter standard deviation or τ (Phillippo 2024), (2) a more informative *half-*
354 *normal*(scale=0.5) prior on the heterogeneity variance (τ^2) proposed by Seide et al (2020) on

355 theoretic grounds, and (3) a more informative empirically-derived *lognormal*(-3.93, 1.51²) prior
356 on τ^2 suggested by Turner et al (2015). LOOcv and Bayesian stacking metrics suggested there
357 was no meaningful difference in predictive accuracy between a network meta-regression model
358 with the weakly informative heterogeneity prior and the same meta-regression model but now
359 with a more informative *lognormal* heterogeneity prior (looic difference = -1.6, se = 1.9) — there
360 was no support for the informative prior suggested by Seide et al (2020). There is a strong
361 empirical basis for the *lognormal* heterogeneity prior (Turner et al 2015), so we proceed with
362 inference based on our network meta-regression model with that prior on the heterogeneity
363 parameter variance.

364

3.3. Weighting Design Specific Effects

366 The predicted design-specific conditional effects relative to the reference design (A) based on
367 the best-fit Bayesian network meta-regression model fitted to the weighting designs network
368 structure (Figure 1) are summarised in Figure 2. This summary display shows both the full
369 posterior distribution of each parameter coupled with key summary metrics (median, highest
370 posterior density intervals) that helps to support a more informative means of communicating
371 parameter uncertainty (van der Bles et al 2019). The six branchline weighting designs B-G are
372 all significantly more effective at reducing expected seabird bycatch than weighting design A —
373 the predicted effects are all significantly < 0 on the log response rate scale. Even for design G
374 we can be > 97% sure that the predicted effect is < 0 (in other words, > 97% of the posterior
375 density shown in Figure 2 is < 0). We can be near 100% sure of this reduced seabird bycatch
376 effect for design E and ca. 99% sure for design F. The largest relative effect is for design F but
377 there is greater uncertainty associated with that effect due to the smaller sample size for design-
378 specific comparisons that included F (see Figure 1, Figure S2). Given this uncertainty, it is
379 apparent that weighting designs E and F appear to have similar efficacy for reducing seabird
380 bycatch in pelagic longline fisheries.

381

3.4. Covariate Effects

383 Figure S3 summarises the conditional effects for the covariates or predictors of seabird bycatch
384 rate included in our best-fit network meta-regression model. The most prominent conditional
385 effects were for (1) the apparent decreased seabird bycatch risk for pelagic longline fisheries
386 operating in the South Pacific region compared to fisheries in the North Atlantic region, and (2)
387 the apparent decreased seabird bycatch risk given the use of night-setting or side-setting as
388 seabird mitigation measures. Note that all covariates were study-specific marginal covariate
389 summaries of aggregate data and not based on set-specific predictors.

390 Despite inclusion of these covariates as main-effects, there remained moderate
391 between-study heterogeneity, by comparison with expected treatment effects (Figure 2, see
392 also Phillippo et al 2020), that was unaccounted for by our best-fit network meta-regression
393 model (posterior mean $\tau = 2.02$; 95% quantile-based CI: 1.67-2.5). High between-study
394 heterogeneity is an expected and common finding for nonexperimental study-based meta-
395 analyses (Turner et al 2015). The between-study heterogeneity apparent in our study might be
396 considerably reduced if we had included covariate interaction terms with say the weighting
397 designs in our models but such interactions were precluded due to unbalanced covariate or
398 effect modifier combinations (Phillippo et al. 2020) — for instance, not all 7 designs were
399 deployed across all 5 ocean regions and so precluded an ocean*design interaction term.

400

3.5. Checking for (In)Consistency

402 The global consistency test compared the predictive accuracy of an unrelated mean effects or
403 UME model (Spinelli 2022) with our best-fit seabird bycatch meta-regression or consistency
404 reference model (see Figure 2). The LOOcv and Bayesian stacking metrics suggested that both
405 models had similar predictive accuracy (looic difference = -0.06, se = 2.6; ca 51% weight of

406 evidence in favour of the best-fit model compared to the UME model). This implies that there
407 was no evidence of global inconsistency and hence our network meta-regression model was fit
408 for inference. Local consistency was evaluated here using the network node-splitting approach
409 and compared the direct and indirect evidence of all pairwise contrasts (node-splits) with
410 sufficient data (15 of the 21 possible contrasts). The direct and indirect posterior density
411 distributions for those 15 node-splits are summarised in Figure S4, which shows no evidence of
412 inconsistency. There was no evidence for network model inconsistency based on either the
413 global or local tests.

414 **3.6. Efficacy Ranking**

415 The median posterior rank for each weighting design based on the posterior samples sourced
416 from our best-fit meta-regression model is summarised in Figure 3. Importantly, there was >
417 97% probability that all weighting designs significantly reduced seabird bycatch compared to
418 reference design A with no weight attached within 5m of the hook. Nonetheless, some weighting
419 designs were significantly more effective at reducing seabird bycatch risk. Designs E and F rank
420 the “best”, although there is considerable uncertainty associated with design F due to small
421 sample size, which is a concern for robust treatment rank scoring (Daly et al 2018, Ades et al
422 2024). An alternative ranking approach uses the model-based cumulative rank order
423 probabilities based on the SUCRA curve shown in Figure S5. The posterior ranking plot (Figure
424 3) and the SUCRA-based ranking plot (Figure S5) show very similar expected weighting design
425 rankings — the only difference being a re-order between designs C and B that are in close
426 proximity in rankings and supports the conclusion that the overall rank ordering is robust (Daly
427 et al 2018). Overall, the two best performing designs (E and F) were found to reduce expected
428 seabird bycatch by *ca.* 89% relative to the reference design: E v A ~ 88.8% (95% HDI: 72-99)
429 and F v A ~ 88.9% (95% HDI: 63-99). The next best performing design D reduced expected
430 seabird bycatch by *ca.* 76% relative to the reference design: D v A ~ 76.1% (95% HDI: 43-97).
431
432

433 **3.7. Relative Effect Contrasts**

434 The relative design-specific effects for all 21 pairwise contrasts between all 7 branchline
435 weighting designs based on the best-fit Bayesian network meta-regression model are
436 summarised in Figure 4 — the first six contrasts repeated the main effects summarised in
437 Figure 2. There was a > 98% probability that all 5 RFMO and ACAP prescribed weighting
438 designs (B-F) had lower expected seabird bycatch rates relative to the reference design A. Even
439 design G (not a prescribed RFMOs or ACAP design) was estimated to reduce the expected
440 seabird bycatch rate by *ca.* 60% (95% HDI: 7-96) compared to the reference design. The
441 contrast estimates in Figure 4 are especially useful for considering design effects that might be
442 more cost-effective or more practical in terms of crew safety when deploying pelagic longline
443 fishing gear. Using the E vs B contrast as an example, there is a > 97% probability that design E
444 is more effective at reducing seabird bycatch than design B. Or there is a > 95% probability that
445 design E is more effective at reducing seabird bycatch than either designs B or C. On the other
446 hand, there is a > 97% probability that design G is more likely to increase seabird bycatch
447 compared to design E or a > 94% probability that G is more likely to increase seabird bycatch
448 compared to design F.
449

450 **4. DISCUSSION**

451 **4.1. Relative Efficacies of Alternative Weighting Designs**

452 Unlike conventional pairwise meta-analysis, NMA enables the simultaneous comparison of
453 multiple treatments or interventions within a coherent modelling framework (Ades et al 2024).
454 NMA approaches achieve this by leveraging direct within-study evidence and indirect across-
455
456

457 studies evidence for all treatments linked in a connected network structure — even if some
458 treatments have not been tested in direct or head-to-head comparisons. The foremost product
459 of any NMA is a league table with rank-ordering of all the treatments to help identify for a
460 specific outcome the relative efficacies of treatments within that network.

461 Our Bayesian multilevel NMA modelling approach synthesized all available evidence to
462 assess the relative efficacies of seven branchline weighting designs for mitigating seabird
463 bycatch in pelagic longline fisheries. The resulting league table identifies the most effective
464 designs. Six treatments had significantly lower expected seabird bycatch rates than the
465 reference design with no weight within 5m of the hook, consistent with the expectation that
466 attaching weights far from the hook results in relatively slow baited hook sink rates, which
467 increases seabird bycatch risk (Brothers et al., 1999; ACAP, 2023). The five RFMO and ACAP
468 prescribed designs had lower expected seabird bycatch rates relative to the reference design,
469 ranging from ca. 66% (design B, 95% HDI:21-95) to ca. 89% (design F, 95% HDI: 63-99) lower
470 expected seabird bycatch rates. There was a > 98% probability that all 5 RFMO and ACAP
471 prescribed weighting designs had lower expected seabird bycatch rates relative to the reference
472 design. The one assessed design not prescribed by RFMOs or ACAP was estimated to reduce
473 the expected seabird bycatch rate by ca. 60% (95% HDI: 7-96) compared to the reference
474 design, and we can be > 97% sure that design G has a lower seabird bycatch rate than the
475 reference design.

476 RFMOs and ACAP prescribe multiple options for branchline weighting designs based on
477 the hypothesis that the alternative designs produce similar seabird catch rates. This is an
478 implicit hypothesis of RFMO seabird measures (IATTC, 2011; ICCAT, 2011; WCPFC, 2018;
479 SIOFA, 2022; IOTC, 2023) and explicit in the ACAP (2023) recommendations. Some of the line
480 weighting designs included in RFMO seabird measures may have been derived from member
481 delegations who ensured that contemporary practices of their fisheries were included as options
482 (e.g., WCPFC, 2006). The ACAP (2023) recommended line weighting design options were
483 defined based on mechanistic studies that identified weighting designs that produce hook sink
484 rates that were predicted to adequately protect seabirds when used in combination with a tori
485 line, and based on findings from three experiments on branchline weighting design effects on
486 seabird interaction rates (ACAP, 2023). Our study findings do not support this hypothesis, as we
487 found that some designs prescribed by RFMOs and ACAP have significantly different seabird
488 catch risks. Of 10 possible pairwise contrasts of the 5 prescribed weighting designs, 4 pairs (E
489 vs. B, E vs. C, F vs. B, and F vs. C) had >92% probabilities of having significantly different
490 seabird bycatch rates. The non-prescribed weighting design included in the NMA models had
491 similar efficacy as two of the prescribed designs. Following the reference treatment, these 3
492 designs had lowest posterior ranks, and had significantly higher seabird catch rates relative to
493 the two designs with lowest seabird catch rates.

494 One weighting design prescribed by all 5 RFMOs with binding seabird management
495 measures (Table S1) of $\geq 98\text{g}$ attached between $>3.5\text{m}$ and $\leq 4\text{m}$ of the hook had too small a
496 sample size for inclusion in the meta-analysis and has an unknown relative efficacy. Discussed
497 in more detail below, the study was limited due to small sample sizes for some prescribed
498 weighting designs by region and number of fisheries. With more certain estimates of the
499 efficacies of alternative prescribed weighting designs, those with relatively low predicted
500 efficacies could be eliminated to augment seabird conservation. Or, for two RFMOs with seabird
501 measures with two lists of methods hypothesized to be relatively more and less effective
502 (IATTC, 2011; WCPFC, 2018), weighting designs could be split into these two lists based on
503 their predicted relative efficacies.

504 Figure 3 (see also Figure S4) shows the estimated relative efficacies of the 7 designs
505 based on an outcome criterion measured as seabird bycatch. Other key performance criteria to
506 holistically evaluate alternative strategies' simulated outcomes across the multiple populations
507 and stocks of exposed commercial and at-risk bycatch species are described in Supplemental

508 Material Section S8. The design-specific contrasts shown in Figure 4 can be evaluated in
509 association with these other management-focused performance criteria. Moreover, the
510 confidence in any guidelines or recommendations based on such branchline weighting design-
511 specific contrasts can then be assessed quantitatively within a Bayesian NMA framework using
512 techniques such as threshold analysis (Phillippo et al 2019). Threshold analysis quantifies how
513 much the evidence would need to change before a management recommendation would be
514 revised, and an assessment could then be made to predict whether the evidence could plausibly
515 change to a degree that exceeds this threshold (Phillippo et al 2019).
516

517 **4.2. Effects of Other Seabird Bycatch Mitigation Methods and Ocean Region**

518 The apparent differences in relative seabird bycatch rates of some of the five bycatch mitigation
519 methods included as model covariates might be partly explained by the effect of crew behavior
520 on the efficacy of some methods. Most (90%) of records in the assembled dataset fit to the NMA
521 models were derived from observer program datasets. Efficacies of tori lines, blue-dyed bait and
522 side setting can be strongly affected by crew behavior. As a result, they might be less effective
523 in practice, during real-world, commercial fishing operations, as determined through analyses of
524 observer and electronic monitoring data, than findings from experiments, despite the latter
525 having a relatively lower risk of error and bias (Gilman and Chaloupka, 2023). Unlike in
526 experimental studies, observational studies do not experimentally manipulate specific variables
527 and control for others, and unlike in observational studies that apply quasi-experimental
528 statistical modelling approaches, observational studies with nominal estimates do not
529 standardize effort and do not explicitly account for simultaneous variability in potentially
530 informative predictors of a response (Cox et al. 2007; Luján and Todt 2021). Contrary to ACAP
531 (2023) criteria for identifying ‘best practice’ seabird bycatch mitigation methods, this is why it is
532 important to validate that the efficacy of an intervention when used under controlled conditions
533 is of similar effectiveness when employed in real-world conditions through observational and
534 ‘pragmatic’ studies (Khorsan and Crawford 2014; Pullin et al. 2021). Variability in tori line design
535 components that affect efficacy, including whether single or paired tori lines were employed
536 (Sato et al., 2013), may also have contributed to the relatively low efficacy of this mitigation
537 method. The very small sample size for HookPods (9 records, see Figure S3) precludes any
538 meaningful assessment of relative efficacy. Furthermore, there may have been confounding
539 effects of other operational as well as environmental predictors of seabird bycatch risk that
540 effected the apparent relative efficacies of these 5 seabird bycatch mitigation methods – for
541 example, all HookPod records also had simultaneous use of night setting. Or, for example, crew
542 might employ a particular seabird bycatch mitigation method, such as a tori line or blue-dyed
543 bait, in areas and periods where high seabird bycatch rates are expected to occur, or in
544 response to observing high seabird interactions or high seabird density.

545 Due to the relationship between seabird density and bycatch risk, the apparent regional
546 differences in relative seabird bycatch risk might be due to regional differences in the local
547 abundance of exposed species of seabirds that are susceptible to pelagic longline capture, and
548 thus regional differences in the magnitude of seabirds scavenging from longline vessels (Gilman
549 et al., 2005). The density of seabirds attending vessels might also affect their scavenging
550 behavior, where up to some threshold, the larger the local seabird abundance, the more intense
551 competitive scavenging behavior and risk of capture will be (Gilman et al., 2005). Regional
552 variability in the composition of seabird species complexes attending vessels might also partly
553 explain this finding. The effect of hierarchical competitiveness between seabird species and
554 between individuals within species, the diving behavior and capabilities of individual species in a
555 complex attending vessels, as well as the occurrence and frequency of secondary interactions,
556 can affect seabird bycatch risk (Jimenez et al., 2012; Melvin et al., 2014). There were also likely
557 regional differences in a wide range of explanatory predictors of seabird bycatch risk that were
558 not able to be accounted for in our NMA, including operational characteristics of fishing methods

559 and gear designs (including regional differences in the use of other seabird bycatch mitigation
560 methods), environmental variables, seabird density and species composition of the assemblage
561 attending vessels, and the spatiotemporal distribution of fishing effort.

562

563 **4.3. Study Limitations**

564 NMA is a powerful statistical procedure to assess multiple treatment effects within a coherent
565 modelling framework (Ades et al 2024). Nonetheless it is important to acknowledge the following
566 limitations when interpreting our network meta-regression model findings. We used the arm-
567 based modelling approach as opposed to using treatment differences in a contrast-based
568 approach (Chu et al 2024) — the arm-based approach is far more flexible and it also supports
569 estimation of pairwise contrast effects (Ades et al 2024, Chu et al 2024). However, one ongoing
570 conceptual concern with arm-based approaches to NMA is that randomisation integrity within
571 trials or studies might not be preserved in the network and that this might bias estimated relative
572 treatment effects — known as the principle of concurrent control (Chu et al 2024) and related to
573 the pattern of missing treatment combinations within a study or trial since not all the trials
574 include all treatments as shown for example in Figure S2. This is an area of ongoing statistical
575 research.

576 Importantly, our study is based on study-specific aggregate or summarised data and not
577 set-specific data from each study. Consequently, all the study-specific covariates were also
578 marginal covariate summaries. This can lead to covariate imbalance across the 7 weight
579 designs and potentially increase the risk of bias in our model-based inference. This is also an
580 area of ongoing statistical research.

581 Moreover, while comprehensive, our study is nonetheless based on a small sample of
582 studies from a small number of fisheries. Sample size limitations prevented us from including
583 various interaction terms in our network model such as ocean region-specific weighting design
584 effects, which might well be important. Furthermore, our NMA model could be mis-specified if
585 the branchline weighting designs were correlated with other potentially informative predictors of
586 seabird bycatch risk that were not available for inclusion in our study. Examples of such
587 potentially informative covariates include: seabird density attending vessels, species
588 composition of the exposed seabird complex attending the fishing vessels, seabird species
589 functional groups such as shallow or deep-diving species, environmental conditions, and other
590 operational characteristics of fishing methods and gear designs (Brothers et al., 1999; Jimenez
591 et al., 2014; Soriano-Redondo et al., 2016; Bi et al., 2020). Designs E and G had broad
592 definitions that could include a wide range of combinations of weight amounts and leader
593 lengths that could have large ranges of efficacies at reducing seabird bycatch risk — with larger
594 sample sizes these categories could be split into narrower design definitions.

595

596 **4.4. Conclusions**

597 Bycatch in pelagic longline fisheries is a major threat to some populations of pelagic seabirds,
598 particularly albatrosses and petrels (Anderson et al. 2011; Phillips et al., 2016). Managing
599 longline weighting designs is one approach for seabird bycatch mitigation that also supports
600 flexible compliance monitoring. Using a Bayesian multilevel network meta-regression modelling
601 approach, our study synthesized available evidence to assess the relative efficacies of
602 alternative branchline weighting designs, including those prescribed by RFMOs and ACAP. All
603 six assessed weighting designs were significantly more effective at reducing seabird bycatch
604 than a reference design of no weight within 5m of the hook, ranging from reductions of ca. 60%
605 (95% HDI: 7-96) to ca. 89% (95% HDI: 63-99) in the estimated seabird bycatch rates. There
606 was a >97% probability that all weighting designs significantly reduced seabird bycatch
607 compared to the reference design. But some RFMO and ACAP prescribed designs had
608 significantly different relative seabird catch rates. To augment seabird conservation, weighting
609 designs with relatively low or uncertain predicted efficacies could be assigned lower preferences

610 or eliminated. When combined with additional key inputs for bycatch management strategy
611 evaluation (Supplemental Material Section S8), these estimates of the relative efficacies of
612 alternative branchline weighting designs support the robust evaluation of alternative bycatch
613 management strategies for regional pelagic longline fisheries.

614
615

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627
628

629 **Supporting Information**

630 This article includes online Supplemental Material.

631
632

633 **Conflict of Interest Statement**

634 The authors declare that they have no competing financial interests or personal relationships
635 that influenced the work reported in this article.

636
637

638 **Data Availability Statement**

639 Data from journal articles and grey literature materials used in this study are available in
640 Supplemental Material Table S2 and the original publications. Table S2 excludes records
641 derived from national observer programs included in the study for pelagic longline fisheries of
642 New Zealand, Taiwan, U.S. Atlantic Ocean, U.S. Pacific Ocean shallow-set swordfish fishery,
643 and U.S. Pacific Ocean deep-set tuna fishery because these records are subject to
644 confidentiality restrictions that prevent the authors from making the observer data publicly
645 available.

646
647

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882

883 **Figure Captions**

884

885 Figure 1. NETWORK GEOMETRY. Diagram showing the 7 (A-G) branchline weighting designs
886 at the nodes connected by 21 direct network interactions or edges between those nodes that
887 were sourced from 18 individual studies. This shows that our network is fully connected and
888 suitable for subsequent network meta-analysis. Node size is proportional to sample size with
889 connecting line (edge) thickness proportional to number of studies in our assembled dataset
890 that included summary data for comparison between any 2 specific designs (nodes). There are
891 three treatment (design) classes of: a reference design A (0 weight attached $\leq 5m$ from hook),
892 designs B through F, which are prescribed by RFMOs or ACAP, and G is a non-prescribed
893 weighting design. Definitions of designs are in the Methods section 2.2 *Branchline Weighting*
894 *Design Treatments*.

895

896 Figure 2. DESIGN-SPECIFIC CONDITIONAL EFFECTS. Summary of the predicted branchline
897 weighting design-specific conditional effects relative to the reference design A (no weight within
898 5m of the hook) based on the best-fit Bayesian network meta-regression model. Coloured
899 polygon shows the posterior density distribution summary (i.e., a wide and thin polygon
900 indicates low precision from a small sample size, and narrow and tall indicates there was high
901 precision and large sample size.), solid dots = median of the posterior density distribution, thick
902 horizontal line below each polygon shows the 80% highest posterior density interval (HDI) for
903 the density distribution while the thin horizontal line is the 95% HDI. Log rate (response) ratios $<$
904 0 = decreased seabird bycatch risk. Although a small fraction of 5 posterior density distributions
905 were > 0 , (salmon shading) all the 6 contrasts with the reference design (A) were significantly $<$
906 0 .

907

908 Figure 3. DESIGN-SPECIFIC RANK-ORDER. Posterior rank summaries for each of the
909 branchline weighting designs (A-G) based on the best-fit network meta-regression model. Solid
910 dot = posterior median rank, thin horizontal line = 95% credible interval summary of the
911 posterior density distribution for that design-specific ranking simulations.

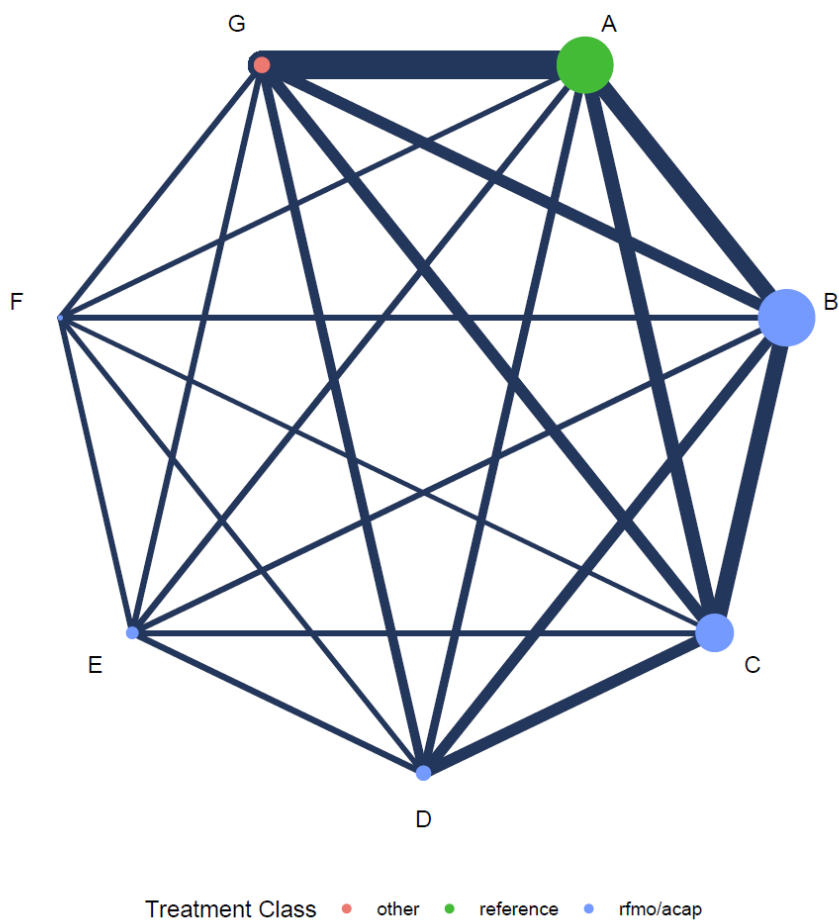
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913 Figure 4. DESIGN-SPECIFIC CONTRASTS. Summary of the predicted log rate ratios for all
914 21 branchline weighting design contrasts based on the best-fit network meta-regression model.
915 Solid dots = conditional median effect estimate, thick horizontal line = 80% credible interval
916 while the thin horizontal line = 95% credible interval. Log rate ratio (response ratio) < 0 =
917 decreased seabird bycatch risk.

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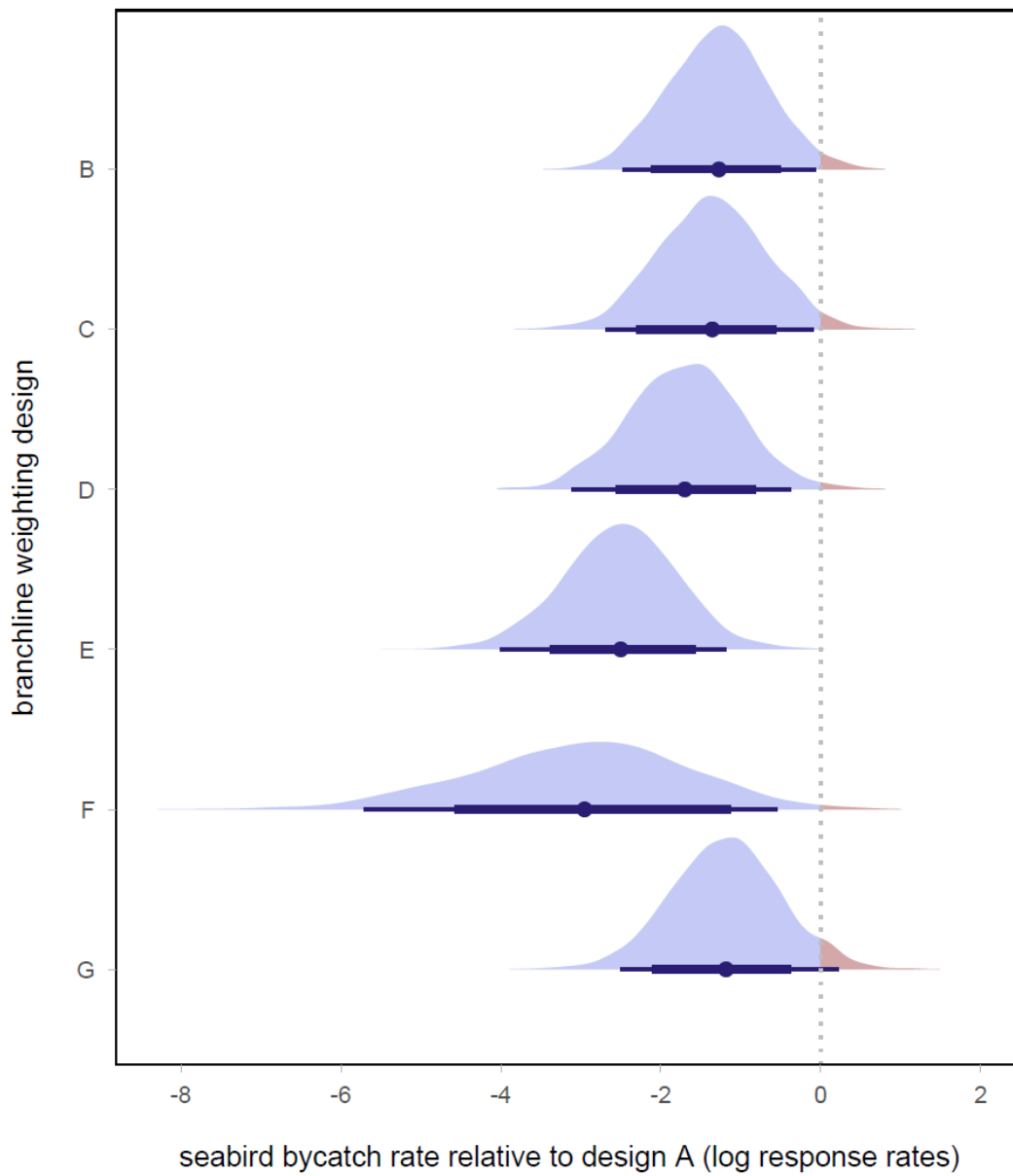
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920 **Figures**
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923 **Figure 1**
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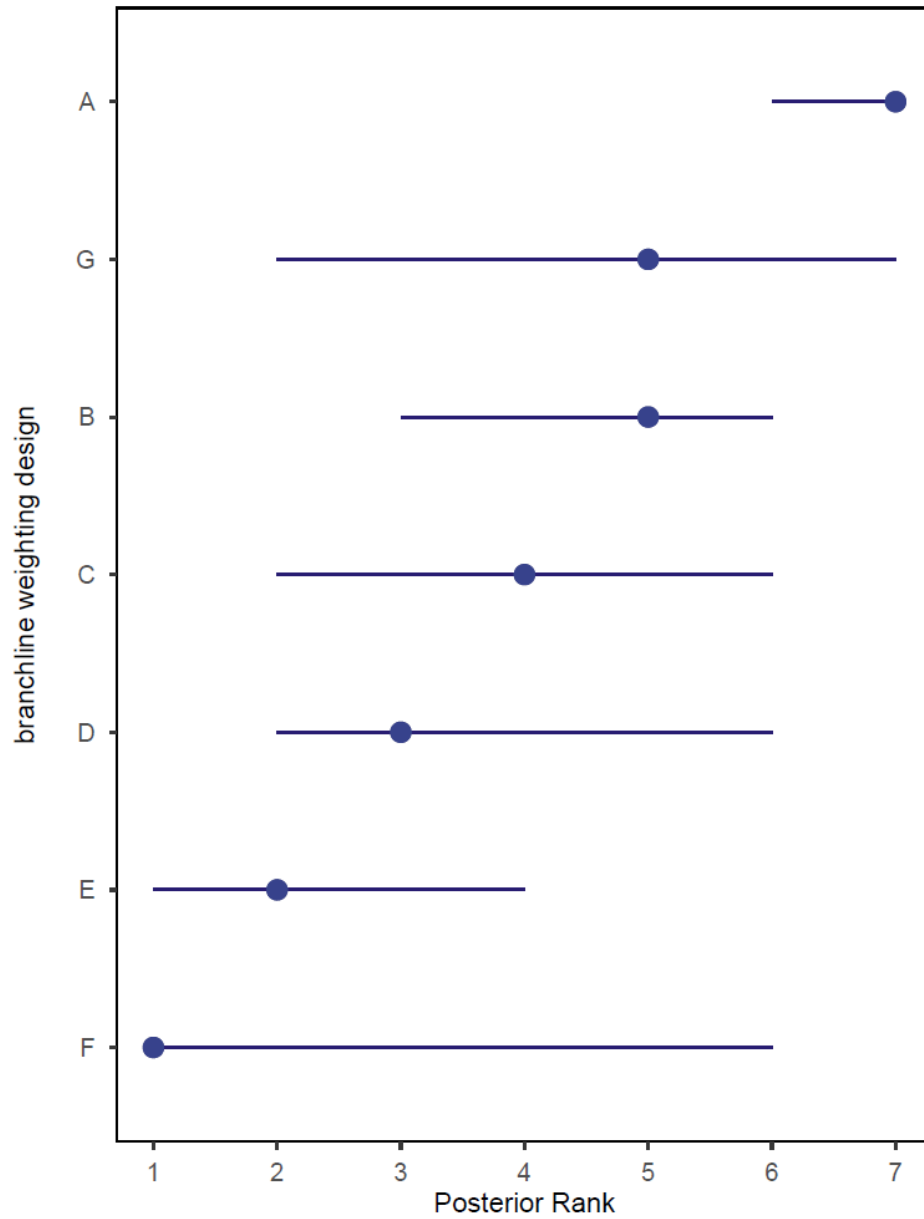
Bayesian ML-NMR predicted branchline weighting effect
posterior density plots (with median and 80% & 95% HDI summaries)



925
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Figure 2

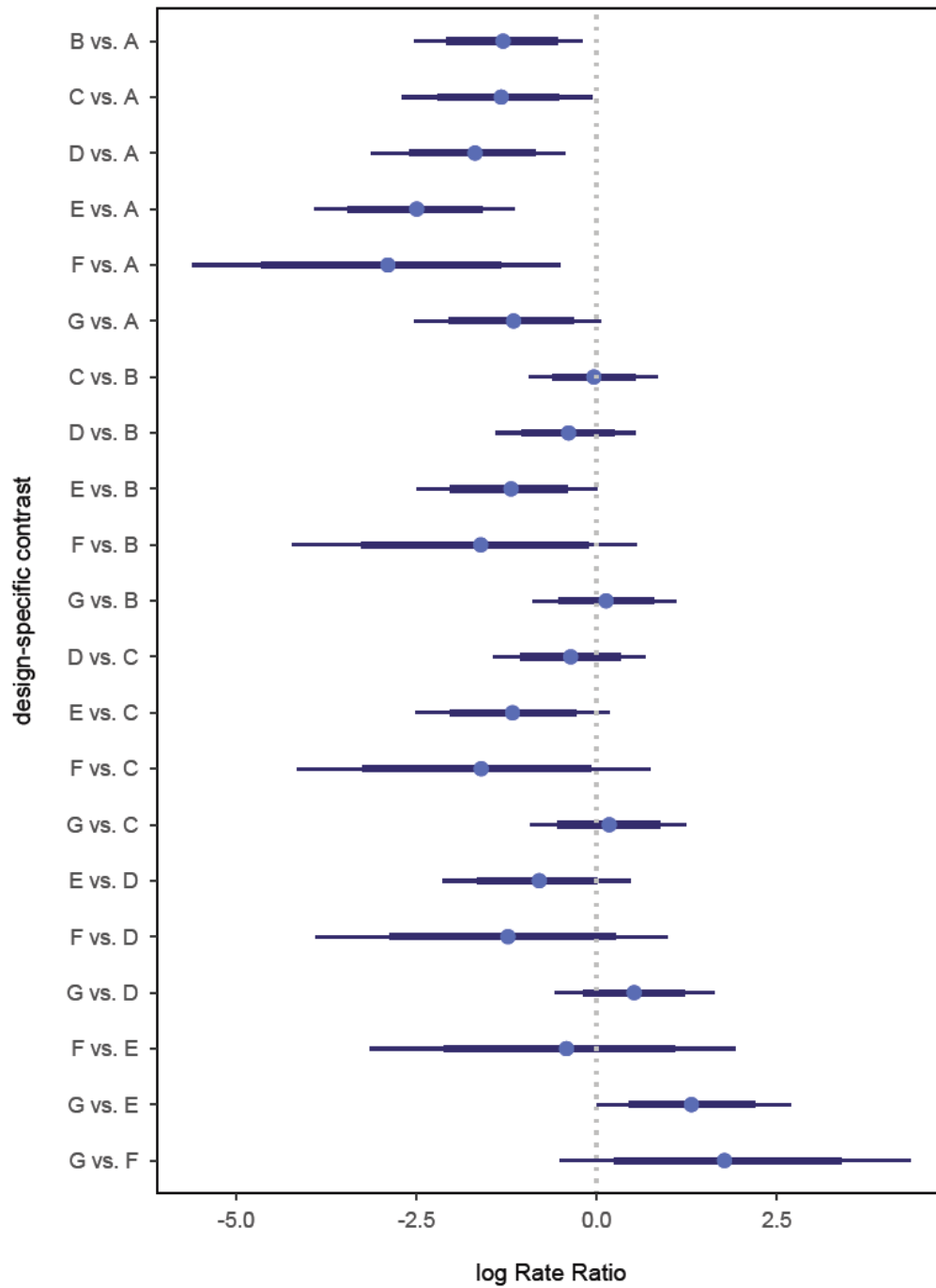
posterior ranking of branchline weighting designs used to mitigate seabird bycatch in pelagic longline fisheries
(median with 95% credible interval summary)



(rank 1 = most effective, 7 = least effective)

928
929 Figure 3
930

relative effects
for the 21 branchline weighting design contrasts
 (median with 80% and 95% credible interval summary)



931
 932 Figure 4